Probabilistic Graphical Models

Lecture 16 – Sampling

CS/CNS/EE 155
Andreas Krause

Announcements

- Homework 3 due today
- Project poster session on Friday December 4 (tentative)
- Final writeup (8 pages NIPS format) due Dec 9

Approximate inference

Three major classes of general-purpose approaches

Message passing

E.g.: Loopy Belief Propagation (today!)

Inference as optimization

- Approximate posterior distribution by simple distribution
- Mean field / structured mean field
- Assumed density filtering / expectation propagation

Sampling based inference

- Importance sampling, particle filtering
- Gibbs sampling, MCMC
- Many other alternatives (often for special cases)

Variational approximation

- Key idea: Approximate posterior with simpler distribution that's as close as possible to P
 - What is a "simple" distribution?
 - What does "as close as possible" mean?
- Simple = efficient inference
 - Typically: factorized (fully independent, chain, tree, ...)
 - Gaussian approximation
- As close as possible = KL divergence

Finding simple approximate distributions

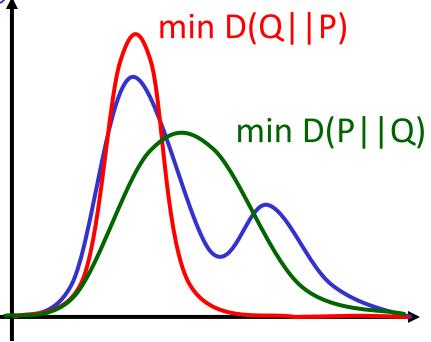
- KL divergence not symmetric; need to choose directions
- P: true distribution; Q: our approximation
- D(P | | Q)

P(x) > 0 => Q(x) > 0

- The "right" way
- Often intractable to compute
- Assumed Density Filtering
- D(Q | | P)

$$P(x) = 0 \Rightarrow Q(x) = 0$$

- The "reverse" way
- Underestimates support (overconfident)
- Mean field approximation
- ullet Both special cases of lpha-divergence



Approximate inference

Three major classes of general-purpose approaches

Message passing

E.g.: Loopy Belief Propagation (today!)

Inference as optimization

- Approximate posterior distribution by simple distribution
- Mean field / structured mean field
- Assumed density filtering / expectation propagation

Sampling based inference

- Importance sampling, particle filtering
- Gibbs sampling, MCMC
- Many other alternatives (often for special cases)

Sampling based inference

- So far: deterministic inference techniques
 - Loopy belief propagation
 - (Structured) mean field approximation
 - Assumed density filtering
- Will now introduce stochastic approximations
 - Algorithms that "randomize" to compute expectations
 - In contrast to the deterministic methods, can sometimes get approximation guarantees
 - More exact, but slower than deterministic variants

Computing expectations

- Often, we're not necessarily interested in computing marginal distributions, but certain expectations:
- Moments (mean, variance, ...)

$$\mathbb{E}_P[X^k] = \int x^k P(x) dx$$

Event probabilities

$$P(\underline{X > c}) = \mathbb{E}_P[\underline{I_{X > c}}] = \int [x > c]P(x)dx$$

Sample approximations of expectations

- \bullet $x_1,...,x_N$ samples from RV X
- Law of large numbers:

$$\mathbb{E}_{P}[f(X)] = \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} f(x_i)$$

- Hereby, the convergence is with probability 1 (almost sure convergence)
- Finite samples: $\approx \frac{1}{\sqrt{2}} \sum_{i=1}^{N} f(x_i)$

How many samples do we need?

Hoeffding inequality
 Suppose f is bounded in [0,C]. Then

$$P\Big(\big| \mathbb{E}_P[f(X)] - \frac{1}{N} \sum_{i=1}^N f(x_i) \big| > \underline{\varepsilon} \Big) \leq 2 \exp(-2N\varepsilon^2/C^2)$$

Thus, probability of error decreases exponentially in N!

$$2\exp(-2N\varepsilon^{2}/C^{2}) < S$$

$$2N\varepsilon^{2}/C^{2} > \log^{\frac{2}{3}}$$

$$N > \frac{1}{\varepsilon^{2}}\cdot C^{2} \cdot \log^{\frac{2}{3}}$$

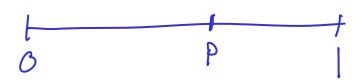
Need to be able to draw samples from P

Sampling from a Bernoulli distribution

X ~ Bernoulli(p)

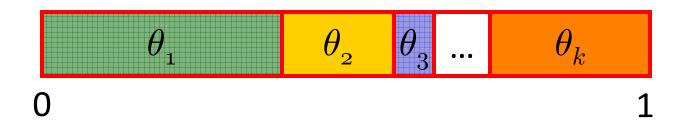
• How can we draw samples from X?

Assame ve can drow uniform distribution [0,1]



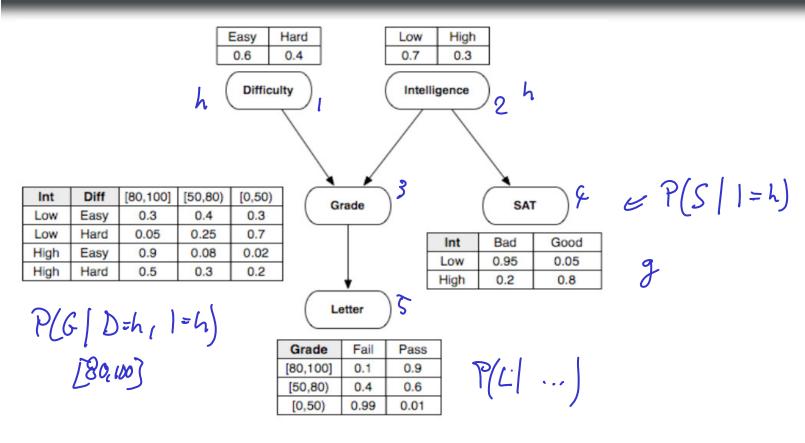
Sampling from a Multinomial

• X ~ Mult($[\theta_1,...,\theta_k]$) where θ_i = P(X=i); $\sum_i \theta_i$ = 1



- Function g: $[0,1] \rightarrow \{1,...,k\}$ assigns state g(x) to each x
- Draw sample from uniform distribution on [0,1]
- Return g⁻¹(x)

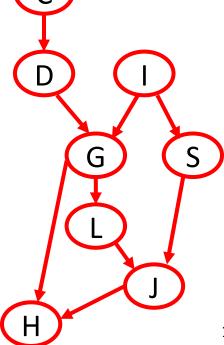
Forward sampling from a BN



Monte Carlo sampling from a BN

- Sort variables in topological ordering X₁,...,X_n
- For i = 1 to n do
 - Sample $x_i \sim P(X_i \mid X_1 = x_1, ..., X_{i-1} = x_{i-1}) = P(X_i \mid \mathcal{Z}_{X_i})$

Works even with high-treewidth models! C

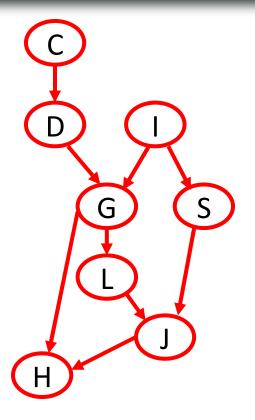


Computing probabilities through sampling

- Want to estimate probabilities
- Draw N samples from BN
- Marginals

$$P(H=y) = \mathbb{E}_{P}[I_{H=y}] = \sum_{x} P(x) \cdot I_{H=y}(x)$$

$$\approx 15^{N} I_{H>y}(x^{(i)}) = \underbrace{(out(H=y))}_{N=1}$$



Conditionals

$$P(D=h|H=m) = \frac{P(D=h,H=m)}{P(H=m)} = \frac{Count(D=h,H>m)}{Count(H=m)}$$

Rejection sampling

Rejection sampling

Collect samples over all variables

$$\widehat{P}(\mathbf{X}_A = \mathbf{x}_A \mid \mathbf{X}_B = \mathbf{x}_B) \approx \frac{Count(\mathbf{x}_A, \mathbf{x}_B)}{Count(\mathbf{x}_B)}$$

- Throw away samples that disagree with x_B
- Can be problematic if $P(X_B = X_B)$ is rare event

Sample complexity for probability estimates

Absolute error:

$$Prob(|\widehat{P}(\mathbf{x}) - P(\mathbf{x})| > \varepsilon) \le 2\exp(-2N\varepsilon^2)$$

Relative error:

$$Prob(\widehat{P}(\mathbf{x}) < (1+\varepsilon)P(\mathbf{x})) \le 2\exp(-NP(\mathbf{x})\varepsilon^2/3)$$

Estimating low probability events

Sampling from rare events

- Estimating conditional probabilities $P(X_A \mid X_B = x_B)$ using rejection sampling is hard!
 - The more observations, the unlikelier $P(X_B = X_B)$ becomes
- Want to directly sample from posterior distribution!

Sampling from intractable distributions

Given unnormalized distribution

$$P(X) \propto Q(X) = P(X_i X_{obs} = X_{obs})$$

- Q(X) efficient to evaluate, but normalizer intractable
- For example, $Q(X) = \prod_j \Psi(C_j)$
- Want to sample from $P(X) = \frac{1}{2}Q(\kappa)$
- Ingenious idea:

Can create Markov chain that is efficient to simulate and that has stationary distribution P(X)

M XA (XB= KB) & P(XA 1XB = XB)

Markov Chains

 A Markov chain is a sequence of RVs, X₁,...,X_N,... with



- Prior P(X₁)
- Transition probabilities P(X_{t+1} | X_t)
- A Markov Chain with $P(X_{t+1} \mid X_t) > 0$ has a unique stationary distribution $\mu(X)$, such that for all $X \mid \lim_{N \to \infty} P(X_N = x) = \mu(x)$

The stationary distribution is independent of $P(X_1)$

Simulating a Markov Chain

Can sample from a Markov chain as from a BN:

- Sample $x_1 \sim P(X_1)$
- Sample $x_2 \sim P(X_2 \mid X_1 = x_1)$
- **...**
- Sample $x_N \sim P(X_N \mid X_{N-1} = x_{N-1})$
- •

• If simulated "sufficiently long", sample X_N is drawn from a distribution "very close" to stationary distribution μ

Markov Chain Monte Carlo

- Given an unnormalized distribution Q(x)
- Want to design a Markov chain with stationary distribution

$$\pi(x) = 1/Z Q(x)$$

Need to specify transition probabilities P(x | x')!

Detailed balance equation

• A Markov Chain satisfies the detailed balance equation for unnormalized distribution Q if for all x, x':

$$Q(x) P(x'|x) = Q(x') P(x | x')$$

In this case, the Markov chain has stationary distribution
 1/Z Q(x)

$$\frac{1}{2}Q(x) = \frac{1}{2}\sum_{x'}P(x'|x)Q(x) = \frac{1}{2}\sum_{x'}Q(x')P(x|x')$$

$$= \sum_{x'}M(x)P(x|x')$$

Designing Markov Chains

- 1) Proposal distribution R(X' | X)
 - Given $X_t = x$, sample "proposal" $x' \sim R(X' \mid X = x)$
 - Performance of algorithm will strongly depend on R
- 2) Acceptance distribution:
 - Suppose $X_t = x$
 - With probability $\alpha = \min\left\{1, \frac{Q(x')R(x\mid x')}{Q(x)R(x'\mid x)}\right\}$ set $\mathbf{X}_{\mathsf{t+1}} = \mathbf{x}'$
 - With probability 1- α , set $X_{t+1} = X$

Theorem [Metropolis, Hastings]: The stationary distribution is Z^{-1} Q(x)

Proof: Markov chain satisfies detailed balance condition!

MCMC for Graphical Models

- Random vector $X=(X_1,...,X_n)$ is high-dimensional
- Need to specify proposal distributions R(x'|x) over such random vectors
 - x4: old state
 - x: proposed state, x' ~ R(X' | X=x)
- Examples

$$- R(X'[X=x]) = R(X')$$

$$- R(X'[X]) \qquad X'_{i} \sim R_{i}(X'[X])$$

$$X'_{i} = X_{i}$$

Gibbs sampling

- Start with initial assignment $\mathbf{x}^{(0)}$ to all variables
- For t = 1 to ∞ do
 - Set $x^{(t)} = x^{(t-1)}$
 - For each variable X_i
 - Set \mathbf{v}_i = values of all $\mathbf{x}^{(t)}$ except \mathbf{x}_i
 - Sample $x^{(t)}_{i}$ from $P(X_{i} | \mathbf{v}_{i})$
- Gibbs sampling satisfies detailed balance equation for P
- Key challenge: Computing conditional distributions P(X_i | v_i)

Computing $P(X_i | v_i)$

$$Q(x) = \prod_{i} Y_{i}(C_{i})$$

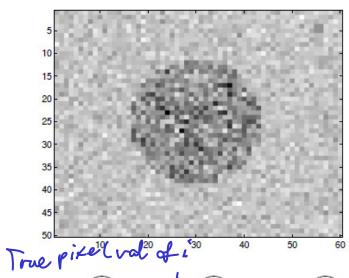
$$P(X_{i} \mid X_{1} ... X_{i-1} \mid X_{i+1} ... X_{m}) = \frac{P(X_{1} ... X_{m})}{P(X_{1} ... X_{m-1} \mid X_{i+1} ... X_{m})}$$

$$= \frac{1}{2} Q(X_{i} ... X_{m}) = \frac{\prod_{i} Y_{i}(C_{i})}{\sum_{X_{i}} \prod_{j} Y_{i}(C_{j})}$$

$$= \frac{1}{2} \frac{1}{2} Q(X_{i} ... X_{m}) = \frac{\prod_{i} Y_{i}(C_{i})}{\sum_{X_{i}} \prod_{j} Y_{i}(C_{j})}$$

$$= \frac{1}{2} \frac{1}{2}$$

Example: (Simple) image segmentation



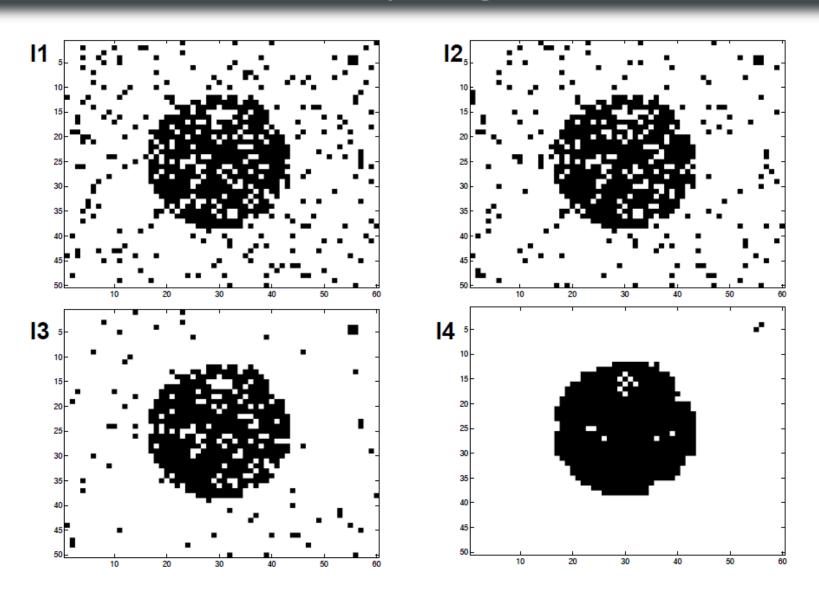
$$P(x) = rac{1}{Z} \prod_i \Phi(x_i) \prod_{(j,k) \in E} \Psi(x_j, x_k)$$

$$\Phi(x_i) = exp\left\{-\frac{(y_i - \mu_{x_i})^2}{2\sigma_{x_i}^2}\right\} \qquad \text{for two pix, with } x_i = x_i$$

$$\Psi(x_i, x_j) = \exp\left\{-\beta(x_i - x_j)^2\right\}$$

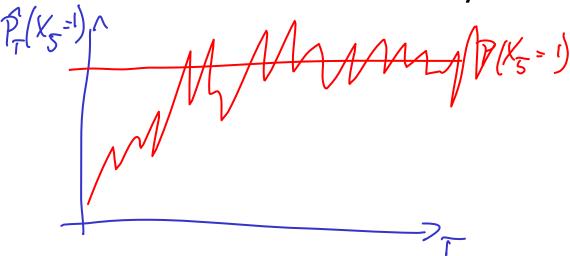
$$v_1$$
 v_2
 v_3
 v_4
 v_4
 v_5
 v_6
 v_7
 v_8
 v_9

Gibbs Sampling iterations



Convergence of Gibbs Sampling

When are we close to stationary distribution?



Summary of Sampling

- Randomized approximate inference for computing expections, (conditional) probabilities, etc.
- Exact in the limit
 - But may need ridiculously many samples
- Can even directly sample from intractable distributions
 - Disguise distribution as stationary distribution of Markov Chain
 - Famous example: Gibbs sampling