

# Probabilistic Graphical Models

## Lecture 5 – Bayesian Learning of Bayesian Networks

CS/CNS/EE 155  
Andreas Krause

# Announcements

- Recitations: Every Tuesday 4-5:30 in 243 Annenberg
- Homework 1 out. Due in class Wed Oct 21
- Project proposals due Monday Oct 19

# Project proposal

- At most 2 pages. One proposal per project
- due Monday Oct 19
- Please clearly specify
  - What is the idea of this project?
  - Who will be on the team?
  - What data will you use? Will you need time "cleaning up" the data?
  - What code will you need to write? What existing code are you planning to use?
  - What references are relevant? Mention 1-3 related papers.
  - What are you planning to accomplish by the Nov 9 milestone?

# Project ideas

- Ideally, do graphical model project related to your research (and, e.g., data that you're working with)
  - Must be a new project started for the class!
- Website has examples for
  - Project ideas
  - Data sets
  - Code

# Project ideas

- All projects should involve using PGMs for some data set, and then doing some experiments
- Learning related
  - Experiment with different algorithms for structure / parameter learning
- Inference related
  - Compare different algorithms for exact or approximate inference
- Algorithmic / decision making
  - Experiment with algorithms for value of information, MAP assignment, ...
- Application related
  - Attempt to answer interesting domain-related question using graphical modeling techniques

# Data sets

- Some cool data sets made available specifically for this course!!  
→ Contact TAs to get access to data.
- Exercise physiological data (collected by John Doyle's group)
  - E.g., do model identification / Bayesian filtering
- Fly data (by Pietro Perona and Michael Dickinson et al.)
  - "Activity recognition" – what are the patterns in fly behavior? Clustering / segmentation of trajectories?
- Urban challenge data (GPS data + LADAR + Vision) by Richard Murray et al.
  - Sensor fusion using DBNs; SLAM
- JPL MER data by Larry Matthies et al.
  - Predict slip based on orbital imagery + GPS tracks
  - Segment images to identify dangerous areas for rover
- LDPC decoding
  - Compare new approximate inference techniques with Loopy-BP
- Other open data sets mentioned on course webpage

# Code

- Libraries for graphical modeling by Intel, Microsoft, ...
- Toolboxes
  - computer vision image manipulations
  - Topic modeling
  - Nonparametric Bayesian modeling (Dirichlet processes / Gaussian processes / ...)

# Learning general BNs

	Known structure	Unknown structure
Fully observable	<u>Easy</u> ✓	hard 2.
Missing data	hard 3. (EM)	very hard (last)



# Algorithm for BN MLE

Given BN structure  $G$

For each variable  $X_i$

$$\text{learn } \hat{\theta}_{X_i | \text{Pa}_i} = \frac{\text{Count}(X_i, \text{Pa}_i)}{\text{Count}(\text{Pa}_i)}$$

$\Rightarrow$  globally maximum likelihood estimate  
for fixed structure  $G$

# Structure learning

- Two main classes of approaches:
- Constraint based
  - Search for P-map (if one exists):
  - Identify PDAG
  - Turn PDAG into BN (using algorithm in reading)
  - **Key problem:** Perform independence tests
- Optimization based ← *coming up!*
  - Define scoring function (e.g., likelihood of data)
  - Think about structure as parameters
  - More common; can solve simple cases exactly

# MLE for structure learning

- For fixed structure, can compute likelihood of data

$$\log P(\mathcal{D} \mid \theta_{\mathcal{G}}, \mathcal{G}) = \sum_{\ell} \sum_i \log P(X_i = x_i^{(\ell)} \mid \mathbf{Pa}_i = \mathbf{pa}_i^{(\ell)}) \quad \theta_{\mathcal{G}} \text{ is } \theta$$

$$\stackrel{\text{MLE}}{=} \sum_i \sum_{x_i} \sum_{\mathbf{pa}_i} \text{Count}(x_i, \mathbf{pa}_i) \log \frac{\hat{p}(x_i, \mathbf{pa}_i)}{\hat{p}(\mathbf{pa}_i)}$$

$$= m \sum_i \sum_{x_i} \sum_{\mathbf{pa}_i} \hat{p}(x_i, \mathbf{pa}_i) \log \frac{\hat{p}(x_i, \mathbf{pa}_i) \hat{p}(x_i)}{\hat{p}(\mathbf{pa}_i) \hat{p}(x_i)}$$

$$= m \sum_i \sum_{x_i} \sum_{\mathbf{pa}_i} \hat{p}(x_i, \mathbf{pa}_i) \log \frac{\hat{p}(x_i, \mathbf{pa}_i)}{\hat{p}(x_i) \hat{p}(\mathbf{pa}_i)} + m \sum_i \sum_{x_i} \sum_{\mathbf{pa}_i} \hat{p}(x_i, \mathbf{pa}_i) \log \hat{p}(x_i)$$

$$= m \sum_i \hat{I}(X_i; \mathbf{Pa}_i) - m \sum_i \hat{H}(X_i)$$

# Decomposable score

- Log-data likelihood

$$\log \hat{P}(\mathcal{D} \mid \theta_{\mathcal{G}}, \mathcal{G}) = m \sum_i \hat{I}(X_i, \mathbf{Pa}_i) - m \sum_i \hat{H}(X_i)$$

*independent of graph structure!*

- MLE score decomposes over families of the BN (nodes + parents)
- $\text{Score}(\mathcal{G} ; \mathcal{D}) = \sum_i \text{FamScore}(X_i \mid \text{Pa}_i; \mathcal{D})$
- Can exploit for computational efficiency!

# Finding the optimal MLE structure

- Log-likelihood score:

$$\text{Score}(\mathcal{G}; \mathcal{D}) = \sum_i \hat{I}(X_i, \mathbf{Pa}_i)$$

- Want  $G^* = \operatorname{argmax}_G \text{Score}(G; \mathcal{D})$
- Lemma:  $G \subseteq G' \Rightarrow \text{Score}(G; \mathcal{D}) \leq \text{Score}(G'; \mathcal{D})$

Complete graph  
maximises log data likelihood!

"Information never hurts"

RV  $X$ ,  $A \subset B$

$$H(X|A) \geq H(X|B)$$

$$I(X; A) = H(X) - H(X|A)$$

$$\Rightarrow I(X; B) \geq I(X; A)$$

# Finding the optimal MLE structure

- Optimal solution for MLE is always the fully connected graph!!! ☹
  - ➔ Non-compact representation; Overfitting!!
- Solutions:
  - Priors over parameters / structures (later)
  - Constraint optimization (e.g., bound #parents)

# Chow-Liu algorithm

- For each pair  $X_i, X_j$  of variables compute

$$\hat{P}(x_i, x_j) = \frac{\text{Count}(x_i, x_j)}{m}$$

- Compute mutual information

$$\hat{I}(X_i, X_j) = \sum_{x_i, x_j} \hat{P}(x_i, x_j) \log \frac{\hat{P}(x_i, x_j)}{\hat{P}(x_i)\hat{P}(x_j)}$$

- Define complete graph with weight of edge  $(X_i, X_j)$  given by the mutual information
- Find maximum spanning tree → skeleton
- Orient the skeleton using breadth-first search

# Today: Bayesian learning

- X Bernoulli variable
- Which is better:
  - Observe 1 H and 2 T  $\hat{\theta} = \frac{1}{3}$
  - Observe 10 H and 20 T  $\hat{\theta} = \frac{1}{3}$
  - Observe 100 H and 200 T  $\hat{\theta} = \frac{1}{3}$
- MLE is same in all three cases
- However, should be much more “confident” about MLE if we have more data
  - ➔ Want to model distributions over parameters



# Bayesian learning

- Make prior assumptions about parameters  $P(\theta)$
- Compute posterior

$$P(\theta | D) = \frac{P(\theta) P(D|\theta)}{P(D)} \propto P(\theta) P(D|\theta)$$

Given data  $D$  want to predict

$$P(x|D) = \int P(\theta|D) \underbrace{P(x|\theta)} d\theta$$

In MLE

$$P(x|D) \approx P(x|\hat{\theta}) \quad \hat{\theta} = \operatorname{argmax}_{\theta} P(D|\theta)$$

# Bayesian Learning for Binomial

$$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta)P(\theta)$$

- Likelihood function:

$$P(\mathcal{D} \mid \theta) = \theta^{m_H} (1 - \theta)^{m_T}$$

- How do we choose prior?
  - Many possible answers...
  - Pragmatic approach: Want computationally “simple” (and still flexible) prior

# Conjugate priors

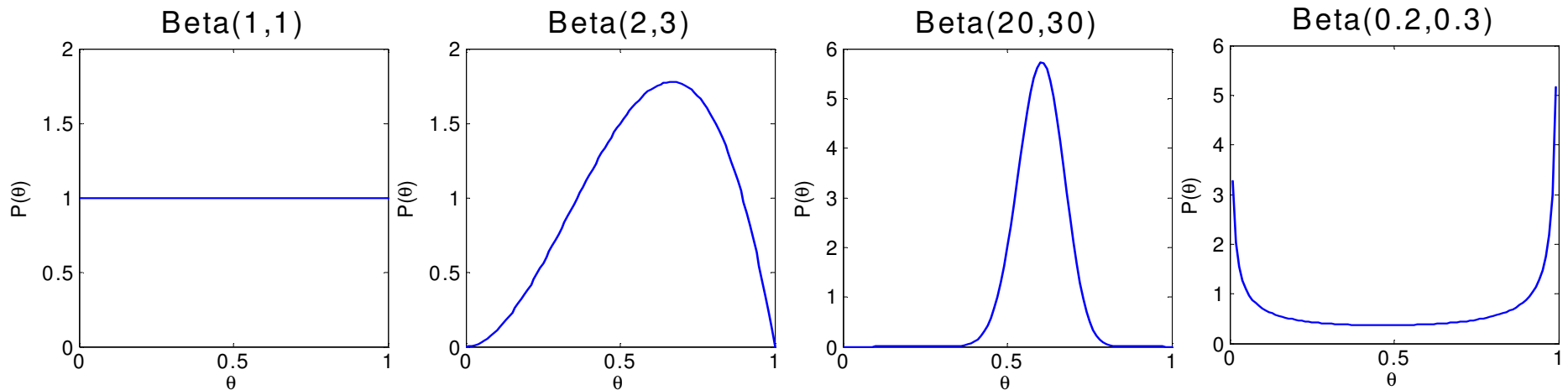
- Consider parametric families of prior distributions:
  - $P(\theta) = f(\theta; \alpha)$
  - $\alpha$  is called “hyperparameters” of prior
- A prior  $P(\theta) = f(\theta; \alpha)$  is called **conjugate** for a likelihood function  $P(D | \theta)$  if  $P(\theta | D) = f(\theta; \alpha')$ 
  - Posterior has same parametric form
  - Hyperparameters are updated based on data  $D$
- Obvious questions (answered later):
  - How to choose hyperparameters??
  - Why limit ourselves to conjugate priors??

# Conjugate prior for Binomial

- Beta distribution

$$\text{Beta}(\theta; \alpha_H, \alpha_T) = \frac{\theta^{\alpha_H-1} (1-\theta)^{\alpha_T-1}}{B(\alpha_H, \alpha_T)}$$

*Normalization constant*



# Posterior for Beta prior

- Beta distribution

$$P(\theta) = \text{Beta}(\theta; \alpha_H, \alpha_T) = \frac{\theta^{\alpha_H-1} (1-\theta)^{\alpha_T-1}}{B(\alpha_H, \alpha_T)}$$

- Likelihood:

$$P(\mathcal{D} \mid \theta) = \theta^{m_H} (1-\theta)^{m_T}$$

- Posterior:

$$P(\theta \mid \mathcal{D}) \propto P(\theta) P(\mathcal{D} \mid \theta) \propto \theta^{\alpha_H+m_H-1} (1-\theta)^{\alpha_T+m_T-1}$$

$$P(\theta \mid \mathcal{D}) = \text{Beta}(\theta; \alpha_H+m_H, \alpha_T+m_T)$$

# Bayesian prediction

- Prior  $P(\theta) = \text{Beta}(\alpha_H, \alpha_T)$       Bernoulli:  $P(X=H) = \theta$
- Suppose we observe  $D = \{m_H \text{ heads, and } m_T \text{ tails}\}$
- What's  $P(X=H \mid D)$ , i.e., prob. that next flip is heads?

$$P(X=H \mid D) = \underbrace{\int \theta P(\theta \mid D) d\theta}_{\text{marginal}} = E[\theta \mid D] = \frac{\alpha_H + m_H}{\alpha_H + \alpha_T + m_H + m_T}$$

# Prior = Smoothing

$$\mathbb{E}[\theta] = \frac{m_H + \alpha_H}{\underbrace{m_H + m_T}_m + \underbrace{\alpha_H + \alpha_T}_{m'}} = \frac{m_H + \gamma m'}{\underbrace{m + m'}_{(*)}}$$

- Where  $m' = \alpha_H + \alpha_T$ , and  $\gamma = \alpha_H / m'$
- $m'$  is called “equivalent sample size”  
 → “hallucinated” coin flips

$$, 0 \leq \gamma \leq 1$$

$$E[\theta] = \frac{m}{m+m'} \underbrace{\frac{m_H}{m}}_{\text{MLE}} + \frac{m'}{m+m'} \underbrace{\gamma}_{\text{prior mean}}$$

$$\begin{array}{lll} m \rightarrow \infty & E[\theta] \rightarrow \text{MLE} & \text{Forget prior} \\ m = 0 & \text{prior} & \end{array}$$

→ Interpolate between MLE and prior mean

# Conjugate for multinomial

- If  $X \in \{1, \dots, k\}$  has  $k$  states:
- Multinomial likelihood

$$P(\mathcal{D} \mid \theta) = \theta_1^{m_1} \theta_2^{m_2} \dots \theta_k^{m_k}$$

where  $\sum_i \theta_i = 1, \theta_i \geq 0$

- Conjugate prior: Dirichlet distribution

$$P(\theta) = \text{Dir}(\theta; \alpha_1, \dots, \alpha_k) = \frac{1}{Z} \prod_i \theta_i^{\alpha_i - 1}$$

- If observe  $D = \{m_1 \text{ 1s}, m_2 \text{ 2s}, \dots, m_k \text{ ks}\}$ , then

$$P(\theta \mid \mathcal{D}) = \text{Dir}(\theta; \alpha_1 + m_1, \dots, \alpha_k + m_k)$$



# Parameter learning for CPDs

- Parameters  $P(X \mid \text{Pa}_X)$
- Have one parameter  $\theta_{X \mid \text{pa}_X}$  for each value of parents  $\text{pa}_X$

$$P(\theta_{X \mid \text{Pa}_X = u}) = \text{Dir}(a_1, \dots, a_r)$$

$$P(\theta_{X \mid \text{Pa}_X = u_1}, \dots, \theta_{X \mid \text{Pa}_X = u_N}) = \prod_u P(\theta_{X \mid \text{Pa}_X = u})$$

"local parameter independence"

# Parameter learning for BNs

- Each CPD  $P(X \mid \text{Pa}_X; \theta_{X|\text{Pa}_X})$  has its own sets of parameters  $P(\theta_{X|\text{Pa}_X})$   
→ Dirichlet distribution
- Want to compute posterior over all parameters

$$P(\theta_{X_1|\mathbf{Pa}_{X_1}}, \dots, \theta_{X_n|\mathbf{Pa}_{X_n}} \mid \mathcal{D})$$

- How can we do this??
- **Crucial assumption:** Prior distribution over parameters factorizes (“<sup>global</sup>parameter independence”)

$$P(\theta_{X_1|\mathbf{Pa}_{X_1}}, \dots, \theta_{X_n|\mathbf{Pa}_{X_n}}) = \prod_i P(\theta_{X_i|\mathbf{Pa}_{X_i}})$$

# Parameter Independence

- Assume

$$P(\theta_{X_1} | \mathbf{Pa}_{X_1}, \dots, \theta_{X_n} | \mathbf{Pa}_{X_n}) = \prod_i P(\theta_{X_i} | \mathbf{Pa}_{X_i})$$

- Why useful?
- If data is fully observed, then

$$P(\theta_{X_1} | \mathbf{Pa}_{X_1}, \dots, \theta_{X_n} | \mathbf{Pa}_{X_n} \mid \mathcal{D}) = \prod_i P(\theta_{X_i} | \mathbf{Pa}_{X_i} \mid \mathcal{D})$$

- I.e., posterior still independent. Why??

# Meta-BN with parameters

Meta-BN

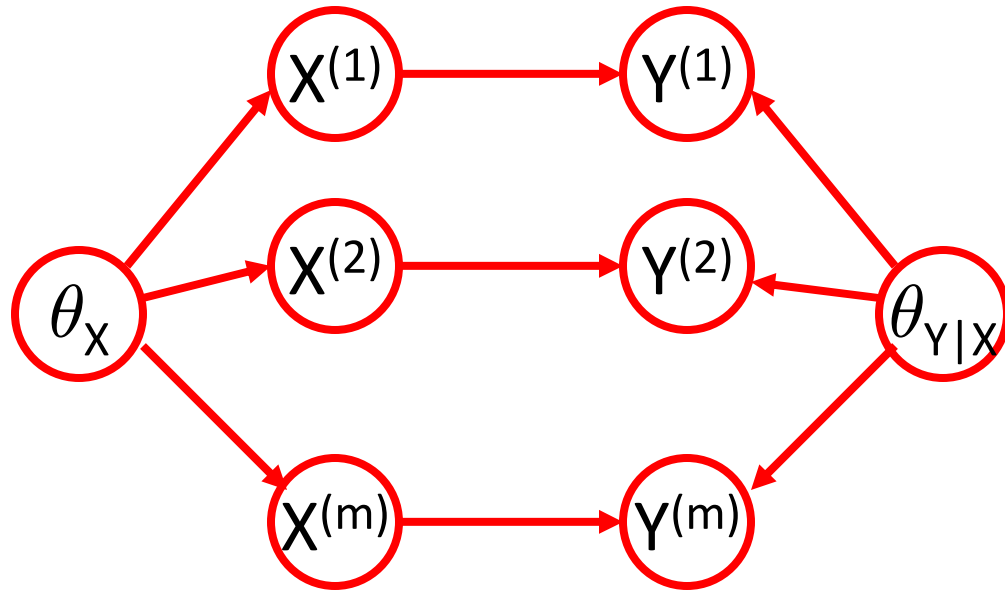
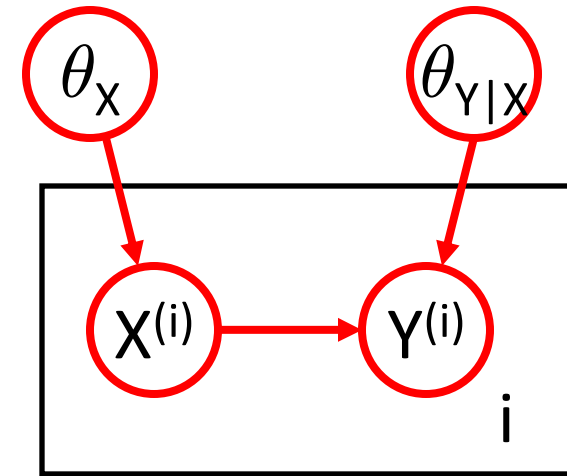


Plate notation



Meta BN contains one copy of original BN per data sample, and one variable for each parameter

Under parameter-independences, data d-separates parameters  
 Also: Parameters d-separate copies of BN:  $P(D, X | \theta) = P(D | \theta) P(X | \theta)$

# Bayesian learning of Bayesian Networks

- Specifying priors helps overfitting
  - Do not commit to fixed parameter estimate, but maintain distribution
- So far: Know how to specify priors over parameters for fixed structure.
- Why should we commit to fixed structure??
- Fully Bayesian inference

$$P(\underline{\mathbf{X}} \mid \mathcal{D}) \propto \sum_{\mathcal{G}} \underbrace{P(\mathcal{G})}_{\text{prior over structures}} \int \underbrace{P(\theta_{\mathcal{G}} \mid \mathcal{G})}_{\text{prior over param.}} \underbrace{P(\mathcal{D} \mid \mathcal{G}, \theta_{\mathcal{G}})}_{\text{likelihood of data}} \underbrace{P(\mathbf{X} \mid \mathcal{D}, \mathcal{G}, \theta_{\mathcal{G}})}_{\substack{\text{likelihood of} \\ \text{pred. vars} \\ = P(\mathbf{X} \mid \mathcal{G}, \theta_{\mathcal{G}})}} d\theta$$

# Fully Bayesian inference

$$P(\mathbf{X} \mid \mathcal{D}) \propto \sum_{\mathcal{G}} P(\mathcal{G}) \int P(\theta_{\mathcal{G}} \mid \mathcal{G}) P(\mathcal{D} \mid \mathcal{G}, \theta_{\mathcal{G}}) P(\mathbf{X} \mid \mathcal{G}, \theta_{\mathcal{G}}) d\theta$$

- $P(\mathcal{G})$ : Prior over graphs

- E.g.:  $P(\mathcal{G}) = \exp(-c \text{Dim}(\mathcal{G}))$

*Dim( $\mathcal{G}$ ) = # free params*

- Called “Bayesian Model Averaging”
- **Hopelessly intractable for larger models**
- Often: want to pick most likely structure:

$$\mathcal{G}^* = \underset{\mathcal{G}}{\operatorname{argmax}} P(\mathcal{G} \mid \mathcal{D}) = \underset{\mathcal{G}}{\operatorname{argmax}} \log P(\mathcal{G}) + \log P(\mathcal{D} \mid \mathcal{G})$$

# Why do priors help overfitting?

$$P(\mathcal{D} \mid \mathcal{G}) = \int P(\mathcal{D} \mid \mathcal{G}, \theta_{\mathcal{G}}) dP(\theta_{\mathcal{G}} \mid \mathcal{G})$$

- This Bayesian Score is tricky to analyze. Instead use:

$$\log \underline{P(\mathcal{D} \mid \mathcal{G})} \approx \log P(\mathcal{D} \mid \mathcal{G}, \widehat{\theta}_{\mathcal{G}}) - \frac{\log m}{2} \text{Dim}(\mathcal{G})$$

- Why??
- **Theorem:** For Dirichlet priors, and for  $m \rightarrow \infty$ :

$$\log P(\mathcal{D} \mid \mathcal{G}) \rightarrow \log P(\mathcal{D} \mid \mathcal{G}, \widehat{\theta}_{\mathcal{G}}) - \frac{\log m}{2} \text{Dim}(\mathcal{G}) + \mathcal{O}(1)$$

# BIC score

$$\log P(\mathcal{D} \mid \mathcal{G}) \approx \underline{\log P(\mathcal{D} \mid \mathcal{G}, \widehat{\theta}_{\mathcal{G}})} - \frac{\log m}{2} \text{Dim}(\mathcal{G})$$

- This approximation is known as **Bayesian Information Criterion** (related to Minimum Description Length)

$$\log P(\mathcal{D} \mid \mathcal{G}) \approx m \sum_i \left( \widehat{I}(X_i; \mathbf{Pa}_i) - \widehat{H}(X_i) \right) - \frac{\log m}{2} \text{Dim}(\mathcal{G})$$

- Trades goodness-of-fit and structure complexity!
- Decomposes along families (computational efficiency!)
- Independent of hyperparameters! (Why??)



# Consistency of BIC

- Suppose true distribution has P-map  $G^*$
- A scoring function  $\text{Score}(G ; D)$  is called **consistent**, if, as  $m \rightarrow \infty$  and probability  $\rightarrow 1$  over  $D$ :
  - $G^*$  maximizes the score
  - All non-I-equivalent structures have strictly lower score
- **Theorem:** BIC Score is consistent!
- Consistency requires  $m \rightarrow \infty$ . For finite samples, priors matter!

# Parameter priors

- How should we choose priors for discrete CPDs?
- Dirichlet (computational reasons). But how do we specify hyperparameters??
- K2 prior:
  - Fix  $\alpha$
  - $P(\theta_{X | Pa_X}) = \text{Dir}(\alpha, \dots, \alpha)$
- Is this a good choice?

$\textcircled{X}$

$\textcircled{Y}$

$$P(\theta_Y) = \text{Dir}(\alpha, \alpha)$$

$\Rightarrow$  Equiv. sample size  
 $2\alpha$

$\textcircled{X} \rightarrow \textcircled{Y}$

$$P(\theta_{Y|X=H}) = \text{Dir}(\alpha, \alpha)$$

$$P(\theta_{Y|X=T}) = \text{Dir}(\alpha, \alpha)$$

$\Rightarrow$  Equiv sample size  
 $4\alpha$

# BDe prior

- Want to ensure “equivalent sample size”  $m'$  is constant
- Idea:
  - Define  $P'(X_1, \dots, X_n)$   
For example:  $P'(X_1, \dots, X_n) = \prod_i \text{Uniform}(\text{Val}(X_i))$
  - Choose equivalent sample size  $m'$
  - Set  $\alpha_{x_i | \text{pa}_i} = m' P'(x_i, \text{pa}_i)$



$$\alpha_y = m' P'(y) = m' \sum_x P'(x, y) = \sum_x \alpha_{y|x}$$

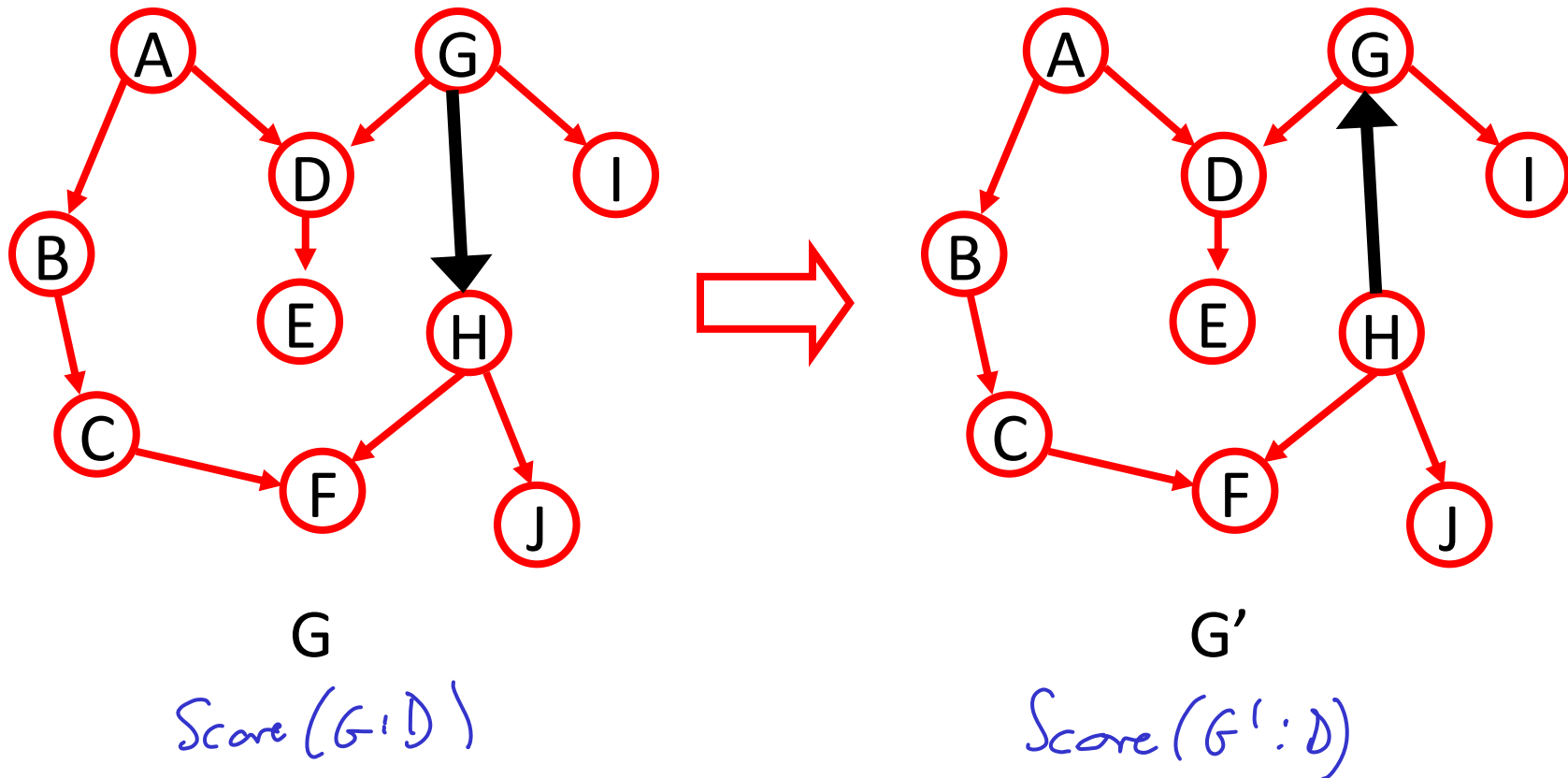
# Bayesian structure search

- Given consistent scoring function  $\text{Score}(G : D)$ , want to find to find graph  $G^*$  that maximizes the score
- Finding the optimal structure is **NP-hard** in most interesting cases (details in reading). ☹️
- Can find optimal tree/forest efficiently (Chow-Liu) 😊
- Want practical algorithm for learning structure of more general graphs..

# Local search algorithms

- Start with empty graph (better: Chow-Liu tree)
- Iteratively modify graph by
  - Edge addition
  - Edge removal
  - Edge reversal
- Need to guarantee acyclicity (can be checked efficiently)
- Be careful with I-equivalence (can search over equivalence classes directly!)
- May want to use simulated annealing to avoid local maxima

# Efficient local search



- Want to avoid recomputing the score after each modification!

# Score decomposability

- Proposition: Suppose we have
  - **Parameter independence**
  - **Parameter modularity**: if  $X$  has same parents in  $G, G'$ , then same prior.
  - **Structure modularity**:  $P(G)$  is product over factors defined over families (e.g.:  $P(G) = \exp(-c|G|)$ )
- Then  $\text{Score}(D : G)$  **decomposes** over the graph:
$$\text{Score}(G ; D) = \sum_i \text{FamScore}(X_i \mid \text{Pa}_i; D)$$
- If  $G'$  results from  $G$  by modifying a single edge, only need to recompute the score of the affected families!!

# What you need to know

- Conjugate priors
  - Beta / Dirichlet
  - Predictions, updating of hyperparameters
- Meta-BN encoding parameters as variables
- Choice of hyperparameters
  - BDe prior
- Decomposability of scores and implications
- Local search



# Tasks

- Read Koller & Friedman Chapter 17.4, 18.3-5
- Project proposal due Monday Oct 19 (contact TAs or instructor to discuss ideas)
- Homework 1 due Wednesday Oct 21